Baker, et al (2025) “Practitioner’s Guide”

2x2 – four averages and three subtractions. You can identify the ATT with only two groups and only two periods if parallel trends holds from the pre period to the post period expressed as E[Y(0)].

* Two groups, two periods
* No anticipation, Parallel trends and untreated comparison group
* DID = ATT
* The challenge though is that with the 2x2, you cannot easily support the parallel trends claim.
* Every 2x2 **always is equal to**:
  + ATT + PT bias term
  + Under NA, all pre-treatment outcomes are Y(0) or don’t have a treatment effect. Which would mean, if you calculated a 2x2 on the pre-treatment period, under NA, those 2x2s equal PT because ATT=0 in the pre-period under NA.

2xT – Event study

* Two groups, but now you have T time periods (T>2).
* No anticipation, parallel trends, untreated comparison group
* Event study coefficients are each 2x2s and therefore each event study coefficient *no matter where it is found in the regression* is always equal to:
  + ATT + PT (PT bias terms).
  + Under NA, ATT=0, so therefore all pre-treatment event study coefficients are *only* equal to PT.
* When you are estimating an event study in a 2xT using OLS, each of the coefficients is a *long difference 2x2*.
  + **Pre-treatment:** When you run a standard OLS “interaction model” of the 2xT where you interact the treatment indicator with calendar time indicators, you have to drop a time period, and therefore all coefficients will be measured relative to that omitted time period. And this is called “long differences”. OLS cannot do it any other way.
  + **Pre-treatment:** But since each coefficient is a 2x2, we could *manually* have calculated pre-treatment coefficients as 2x2s but used the neighbor, not a fixed baseline, in which case you have a *rolling comparison*. This called sometimes “short gaps”. OLS cannot do this – we can do it manually, but OLS is not manual.
  + **Post-treatment:** Everyone, no matter the method, cannot calculate those post-treatment event study coefficients any other way than using a fixed baseline. So all post-treatment coefficients in an event study are always calculated using *fixed dropped omitted baseline time period*. Because it’s literally impossible to use “the neighbor” because the neighbor is treated and we need no anticipation (an untreated comparison).

2x2xT

GxT

DxT

Unconditional parallel trends

Males earnings grows +2 a year

Female earnings grows +1 a year

* Randomly assign the treatment to this population. What is the average earnings growth in the treatment group versus the control group.
* If it was truly random, then the treatment is assigned independent of Y(0) and it is also assigned independent of *changes* in Y(0) which would mean that the average change in Y(0) is the same in the treatment as it is in the control group because the covariates that cause Y(0) trends are balanced.
* Treatment group will be half male; control group will be half female.
* So what is the average change in Y(0) in the treatment group:
  + Treatment group: Change in E[Y(0)] = ( 0.5 x 2) + (0.5 x 1) = 1.5
  + Control group: Change in E[Y(0)] = ( 0.5 x 2) + (0.5 x 1) = 1.5

What if it was not an RCT? What if your observational study was like this:

* Treatment group: 75% male, 25% female
* Control group: 25% male, 75% female

What is the average change in Y(0) for the treatment group vs control group?

* Treatment group: (0.75\*2) + (0.25\*1) = 1.75. On average, untreated potential outcomes in the treatment group grow by 1.75 a year.
* Control group: (0.25\*2) + (0.75\*1) = 1.25. On average, in the control group, untreated potential outcomes grow by 1.25 a year.

THEREFORE, UNCONDITIONAL PARALLEL TRENDS DOES NOT HOLD**. And it does not hold because the covariates that cause trends in untreated potential outcomes are *imbalanced across the two groups***.

**Abadie IPW (inverse probability weighting)**

Weighted 2x2=

E [ ( **long difference**/average number of units treated ) x (D-p(x))/(1-p(x)) ]

**Long difference is “after minus before” average outcomes**.

Abadie IPW 2x2 = w1 x long diff + w0 x long diff

= ( **w1**(Y\_post, D=1) - w1 (Y\_Pre, D=1) ) - ( w0(Y0\_post, D=0) – w0 Y(\_Pre, D=0) )

What is **w1**? That is the weight on treated units (D=1).

W1 = 1/Pr(D=1)

What is **w0?** That is the weight on control units (D=0).

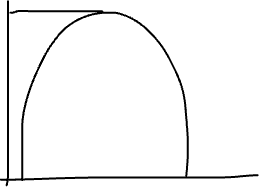
W0 = -p(x)/(1-p(x))

DID in R does ipw, double robust and outcome regression. BUT it also “trims”. If the propensity score is above 0.995, it drops those units.

CSDID does not do that.

In our application, we had **three counties** out of **3,000 counties** with propensity scores above 0.995. And just those 3 counties caused CSDID to be 5 times larger than the R package.

0.25



0.5 0.5 Dbar

OLS weights in the TWFE specification on the 2x2s are holding all else constant **maximized** at Dbar=0.5. What does that mean – any treatment group treated in the **middle of the panel** will get the largest variance weights.

**TWFE = VWATT\_k,Post + VWPT – Delta ATT\_comparison groups**

Example 1:

Assume VWATT = +10

Assume VWPT = 0

Assume comparison group’s change in ATT across all those 2x2s equals +2 (maye it goes from 8 to 10, or 100 to 102).

Then if you estimated the TWFE model – year fixed effects, state fixed effects, no covariates, treatment additively in – then the TWFE coefficient would numerically equal:

**TWFE** = 10 + 0 – 2 = **8**You’d be biased. TWFE is “attenuated” because of the negative weight (-1) on those dynamics. Causes the positive Delta to “flip signs” and pull down the TWFE coefficient away from the VWATT.

Of all the ways that TWFE is biased, this is probably the best way. Because at least the sign is the right sign.

**TWFE = VWATT\_k,Post + VWPT – Delta ATT\_comparison groups**

Example 2:

VWATT = +10

VWPT = 0

Delta ATT for comparison groups = 11 (maybe it went from 2 to 13 or 5 to 16).

What will you get when you run TWFE?

TWFE = +10 + 0 **-11 = -1**

You run TWFE, you’ll get a -1 and in reality the truth is a positive 10. You not only got the answer wrong; you got the sign wrong.

It would be like “I do an experiment and my experiment says if I give a person this drug, they will live 2 more days, but in reality it was they will die immediately”.

Borusyak and Jaravell (2017 or 2018) in the early working paper said “OLS does not satisfy a *no sign flip property*” which is in econometrics considered to be the **bare minimum**.

When you run a regression on a variable that has classical measurement error, the coefficient is **biased towards zero**, but it never reaches zero.